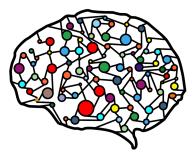
Lecture 28 – Review problems



DSC 40A, Fall 2022 @ UC San Diego

Dr. Truong Son Hy, with help from many others

Announcements

- Final exam tomorrow!
- Review the solutions to previous homeworks and groupworks.
- Identify which concepts are still iffy. Re-watch lecture and ask questions (now!).
- Look at the past exams at https://dsc40a.com/resources.
- Study in groups.
- Make a "cheat sheet".
- Bring a calculator.
- Remember to submit The Course and Professor Evaluations (CAPE) – deadline tonight. If everyone submits CAPE, everyone will get a bonus percentage!

Final schedule

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		FI	12/03/2022	S	7:00p-9:59p	CSB	001				
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	88109			MWF M	4:00p-4:50p 6:00p-6:50p	PCYNH 🖗 PCYNH 🖑	122 122		Resources 19	Evaluat 115	

Time: December 3rd, 2022 – 7:00pm to 10:00pm (3 hours) **Location:** CSB building – room 001 (for my section)

Review problems

Example: Clustering and combinatorics

- Suppose we have a dataset of 15 points, each with two features (x_1, x_2) . In the dataset, there exist 3 "natural" clusters, each of which contain 5 data points.
- Recall that in the k-Means Clustering algorithm, we initialize k centroids by choosing k points at random from our dataset. Suppose k = 3.

Questions:

- 1. What's the probability that all three initial centroids are initialized in the same natural cluster?
- 2. What's the probability that all three initial centroids are initialized in different natural clusters?

Let S denote the whole 2-dimensional space (i.e. sample space). Let E_1 , E_2 , and E_3 denote the clusters (i.e. partitions):

$$E_1 \cap E_2 = \emptyset, \quad E_2 \cap E_3 = \emptyset, \quad E_3 \cap E_1 = \emptyset$$

 $E_1 \cup E_2 \cup E_3 = S.$

Let $\mu_1, \mu_2, \mu_3 \in S$ denote the initial centroids.

For simplicity, we assume S is bounded (i.e. S is the convex hull of all $\{(x_i, y_i)\}_{i=1}^{15}$). We have:

$$P(\mu_i \in E_i) = \frac{|E_i|}{|S|}.$$

1. What's the probability that all three initial centroids are initialized in the same natural cluster? We further assume **uniform** distribution: $P(\mu_i \in E_i) = 1/3$. Because of **independence** in sampling, the final result is:

 $P(\mu_1 \in E_1, \mu_2 \in E_2, \mu_3 \in E_3) = P(\mu_1 \in E_1)P(\mu_2 \in E_2)P(\mu_3 \in E_3) = \frac{1}{27}$

2. What's the probability that all three initial centroids are initialized in different natural clusters?

 $P(\mu_1 \notin E_1, \mu_2 \notin E_2, \mu_3 \notin E_3) = P(\mu_1 \notin E_1)P(\mu_2 \notin E_2)P(\mu_3 \notin E_3) = \frac{8}{27}$

Example: basketball

Suppose we have 6 basketball players who want to organize themselves into 3 basketball teams of 2 players each. Suppose

we have three teams, "Team USA", "Team China", and "Team Lithuania". How many ways can these teams be formed?

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Using permutations:

$$\frac{6!}{2!2!2!} = 90$$

Using combinations:

$$\binom{6}{2} \cdot \binom{4}{2} \cdot \binom{2}{2} = \frac{6!}{4!2!} \cdot \frac{4!}{2!2!} \cdot 1 = \frac{6!}{2!2!2!} = 90$$

Example: basketball, again

Suppose we have 6 basketball players who want to organize themselves into 3 basketball teams of 2 players each. Now,

suppose the teams are irrelevant, and all we care about is the unique pairings themselves. How many ways can these 6 players be split into 3 teams?

Example: basketball, again

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suppose the teams are irrelevant, and all we care about is the unique pairings themselves. How many ways can these 6 players be split into 3 teams?

If the order of 3 teams matters, we have 90 ways to form the teams. If the order does not matter, we divide by 3!. Finally, we get:

 $\frac{90}{3!} = 15$

Example: Lottery

When you buy a Powerball ticket, you select 5 different white numbers from among the numbers 1 through 59 (order of the selection does not matter), and one red number from among the numbers 1 through 35. How many different Powerball tickets can you buy?

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If you check out the Powerball web site you will see that you need to select 5 distinct white numbers, so you can do this $\binom{59}{5} = 5,006,386$ ways. Then you can pick the red number $\binom{35}{1} = 35$ ways so the total number of tickets is:

 $\binom{59}{5} \cdot \binom{35}{1} = 5,006,386 \cdot 35 = 175,223,510.$

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(a) How many (different) samples (of size 4) are possible?

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(c) How many samples have 2 red and 2 white marbles?

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 $\binom{10}{4} = 210$

(c) How many samples have 2 red and 2 white marbles?

$$\binom{10}{2} \cdot \binom{5}{2} = 45 \cdot 10 = 450$$

(d) How many samples (of size 4) have exactly 3 red marbles?

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$$\binom{10}{3} \cdot \binom{5}{1} = 120 \cdot 5 = 600$$

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(d) How many samples (of size 4) have exactly 3 red marbles?

$$\binom{10}{3} \cdot \binom{5}{1} = 120 \cdot 5 = 600$$

(e) How many samples (of size 4) have at least 3 red? The answer is the number of samples with 3 red plus the number of samples with 4 red:

$$\binom{10}{3} \cdot \binom{5}{1} + \binom{10}{4} \cdot \binom{5}{0} = 120 \cdot 5 + 210 \cdot 1 = 810$$

(f) How many samples (of size 4) contain at least one red marble?

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One answer is "the number with exactly 1" + "the number with exactly 2" + "the number with exactly 3" + "the number with exactly 4":

$$\binom{10}{1} \cdot \binom{5}{3} + \binom{10}{2} \cdot \binom{5}{2} + \binom{10}{3} \cdot \binom{5}{1} + \binom{10}{4} \cdot \binom{5}{0}$$

= 10.10+45.10+120.5+210.1 = 100+450+600+210 = 1,360 But it is a slow computation!

(f) How many samples (of size 4) contain at least one red marble?

One answer is "the number with exactly 1" + "the number with exactly 2" + "the number with exactly 3" + "the number with exactly 4":

$$\binom{10}{1} \cdot \binom{5}{3} + \binom{10}{2} \cdot \binom{5}{2} + \binom{10}{3} \cdot \binom{5}{1} + \binom{10}{4} \cdot \binom{5}{0}$$

 $= 10 \cdot 10 + 45 \cdot 10 + 120 \cdot 5 + 210 \cdot 1 = 100 + 450 + 600 + 210 = 1,360$

But it is a slow computation! The faster answer is the total number of samples minus the number of samples with no red marbles:

$$\binom{15}{4} - \binom{10}{0} \cdot \binom{5}{4} = 1,365 - 5 = 1,360.$$

Example: bitstrings

What is the probability of a randomly generated bitstring of length 5 having the same first two bits? Assume that each bit is equally likely to be a 0 or a 1.

We can start the bitstring as 00 or 11. The rest of the string does not matter. Because, there are 4 different ways to start the bitstring: {00, 01, 10, 11}. The final result is:

Example: bitstrings, again

What is the probability of a randomly generated bitstring of length 5 having the same first two bits, if we know that the bitstring has exactly four 0s? Assume that each bit is equally likely to be a 0 or a 1.

The first two bits must be two zeros, because there are exactly four 0s. There is also exactly one bit 1. We can only put bit 1 among the 3rd, 4th and 5th position. Thus, only 3 possibilities:

{00100,00010,00001}

 $\frac{3}{32}$.

The final result is:

Given a 2 × 2 matrix:

$$X = \begin{pmatrix} x_{11} & x_{12} \\ x_{21} & x_{22} \end{pmatrix},$$

its inverse (if exists) is given by:

$$X^{-1} = \frac{1}{\det(X)} \cdot \begin{pmatrix} x_{22} & -x_{12} \\ -x_{21} & x_{11} \end{pmatrix},$$

where det(X) is the determinant of *X*:

$$\det(X) = x_{11}x_{22} - x_{21}x_{12}.$$

Given:

$$X = \begin{pmatrix} 5 & 2 \\ -7 & -3 \end{pmatrix},$$

what is X^{-1} ?

We have the determinant:

$$\det(X) = 5 \cdot (-3) - (-7) \cdot 2 = -1,$$

and the inverse:

$$X^{-1} = \frac{1}{-1} \cdot \begin{pmatrix} -3 & -2 \\ 7 & 5 \end{pmatrix} = \begin{pmatrix} 3 & 2 \\ -7 & -5 \end{pmatrix}.$$

Easy to verify:

$$\begin{pmatrix} 5 & 2 \\ -7 & -3 \end{pmatrix} \begin{pmatrix} 3 & 2 \\ -7 & -5 \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$
 and $\begin{pmatrix} 3 & 2 \\ -7 & -5 \end{pmatrix} \begin{pmatrix} 5 & 2 \\ -7 & -3 \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$.

Given:

$$X = \begin{pmatrix} 5 & 2 \\ 5 & 2 \end{pmatrix},$$

what is X^{-1} ?

We have the determinant:

$$\det(X)=5\cdot 2-5\cdot 2=0,$$

thus there is no inverse for X.

Given:

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We have the determinant:

$$\det(X) = 5 \cdot (-7) - (-7) \cdot 5 = 0,$$

thus there is no inverse for X.

Example: Linear regression by pseudo-inverse

Apply linear regression in matrix form of this data

 $D=\{(0,0),(1,1)\}.$

Fundamentally, we fit a line.

We have:

$$X = \begin{pmatrix} 1 & 0 \\ 1 & 1 \end{pmatrix} \text{ and } \vec{y} = \begin{pmatrix} 0 \\ 1 \end{pmatrix}.$$

The closed-form formula for linear regression is:

 $\vec{w} = (X^T X)^{-1} X^T \vec{y}.$

Example: Linear regression by pseudo-inverse

First, we compute the covariance matrix:

$$X^{\mathsf{T}}X = \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} 1 & 0 \\ 1 & 1 \end{pmatrix} = \begin{pmatrix} 2 & 1 \\ 1 & 1 \end{pmatrix}.$$

We have $det(X^T X) = 2 \cdot 1 - 1 \cdot 1 = 1$, and:

$$(X^T X)^{-1} = \frac{1}{1} \begin{pmatrix} 1 & -1 \\ -1 & 2 \end{pmatrix} = \begin{pmatrix} 1 & -1 \\ -1 & 2 \end{pmatrix}.$$

The pseudo-inverse of X is:

$$(X^T X)^{-1} X^T = \begin{pmatrix} 1 & -1 \\ -1 & 2 \end{pmatrix} \cdot \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ -1 & 1 \end{pmatrix}.$$

The final result for \vec{w} is:

$$\vec{w} = \begin{pmatrix} w_0 \\ w_1 \end{pmatrix} = (X^T X)^{-1} X^T \vec{y} = \begin{pmatrix} 1 & 0 \\ -1 & 1 \end{pmatrix} \cdot \begin{pmatrix} 0 \\ 1 \end{pmatrix} = \begin{pmatrix} 0 \\ 1 \end{pmatrix}.$$

We have the sigmoid function:

$$\sigma(x) = \frac{1}{1 + e^{-x}}$$

Let's derive its' derivative!

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$$\frac{d}{dx}\sigma(x) = \frac{d}{dx}\frac{1}{1+e^{-x}} = \frac{d}{dx}(1+e^{-x})^{-1} = -(1+e^{-x})^{-2}\frac{d}{dx}(1+e^{-x})$$
$$\Leftrightarrow \frac{d}{dx}\sigma(x) = (1+e^{-x})^{-2}e^{-x} = \frac{e^{-x}}{(1+e^{-x})^2} = \frac{1}{1+e^{-x}} \cdot \frac{e^{-x}}{1+e^{-x}}$$
$$\Leftrightarrow \frac{d}{dx}\sigma(x) = \sigma(x) \cdot \frac{e^{-x}+1-1}{1+e^{-x}} = \sigma(x) \cdot [1-\sigma(x)].$$

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Show that $\sigma(x)$ is a monotonically increasing function and $0 < \sigma(x) < 1$.

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Show that $\sigma(x)$ is a monotonically increasing function and $0 < \sigma(x) < 1$.

We have:

$$\sigma'(x)=\frac{e^{-x}}{(1+e^{-x})^2}>0,\quad\forall x\in(-\infty,+\infty),$$

because $e^{-x} > 0$ and $(1 + e^{-x})^2 > 0$. Thus $\sigma(x)$ is a monotonically increasing function.

We have $1 + e^{-x} > 0$, thus $\sigma(x) > 0$. Furthermore:

 $\sigma(x) < 1 \Leftrightarrow 1 < 1 + e^{-x} \Leftrightarrow e^{-x} > 0$

holds. Since:

$$\lim_{x \to +\infty} e^{-x} = 0, \quad \lim_{x \to -\infty} e^{-x} = +\infty,$$
$$\lim_{x \to -\infty} \sigma(x) = 0,$$

we get:

 $\lim_{x\to+\infty}\sigma(x)=1.$

Example: Logistic

Given a hypothesis

$$h(\vec{x};\vec{w},w_0)=\sigma(\vec{x}\cdot\vec{w}+w_0).$$

In logistic regression, we have:

$$p(y = 1 | \vec{x}; \vec{w}, w_0) = \sigma(\vec{x} \cdot \vec{w} + w_0).$$

Let's derive the partial derivative of h with respect to \vec{w} and $w_0!$

Example: Logistic

First of all, we have:

$$\frac{d}{dw_0}(\vec{x}\cdot\vec{w}+w_0) = 1,$$
$$\frac{\partial}{\partial\vec{w}}(\vec{x}\cdot\vec{w}+w_0) = \vec{x}.$$

For convenience, let $\bar{h} = \vec{x} \cdot \vec{w} + w_0$. By the chain rule, we get:

$$\frac{dh}{dw_0} = \sigma'(\bar{h}) \cdot \frac{d\bar{h}}{dw_0} = \sigma(\bar{h}) \cdot [1 - \sigma(\bar{h})],$$
$$\frac{\partial h}{\partial \vec{w}} = \sigma'(\bar{h}) \cdot \frac{\partial \bar{h}}{\partial \vec{w}} = \sigma(\bar{h}) \cdot [1 - \sigma(\bar{h})] \cdot \vec{x}.$$