

Lecture 11

Regression and Linear Algebra

DSC 40A, Fall 2024

Announcements

- Homework 3 is due on **Friday, October 25th**.
- Homework 1 scores are available on Gradescope.
 - Regrade requests are due tonight.
- The Midterm Exam is on **Monday, Nov 4th in class**.

Agenda

- Regression and linear algebra.
- Finding the optimal parameter vector
 - by minimizing the projection error (linear algebra).
 - by minimizing empirical risk (multivariate calculus).

Question 🤔

Answer at q.dsc40a.com

Remember, you can always ask questions at q.dsc40a.com!

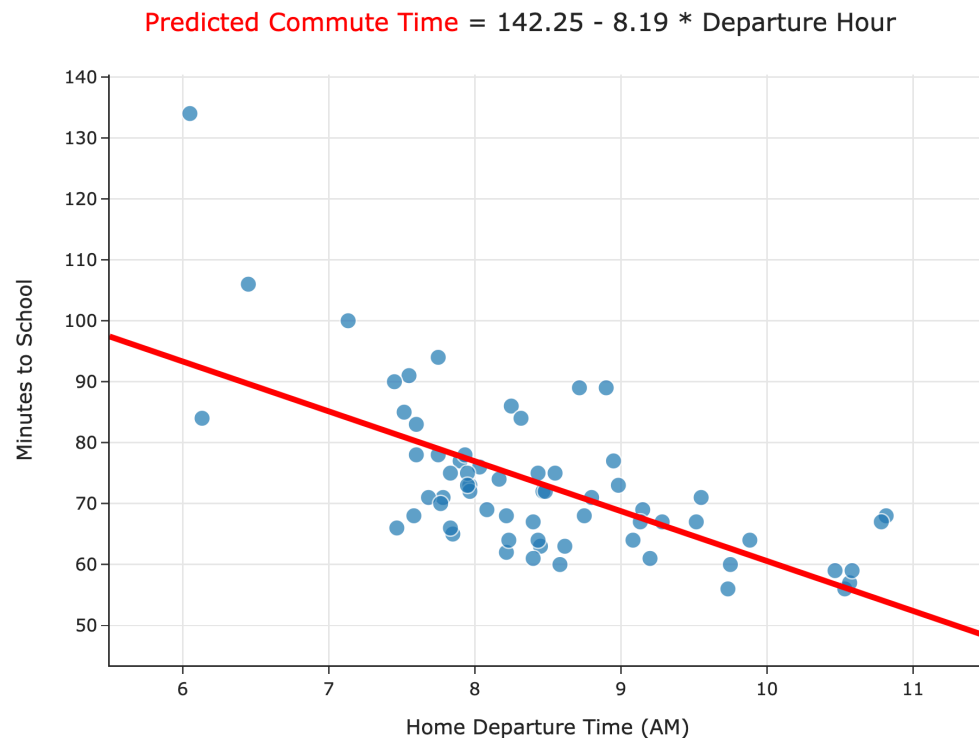
If the direct link doesn't work, click the "🤔 Lecture Questions"
link in the top right corner of dsc40a.com.

Regression and linear algebra

Wait... why do we need linear algebra?

- We want to make predictions using more than one feature.
 - Example: Predicting commute times using departure hour and temperature.
- Thinking about linear regression in terms of **matrices and vectors** will allow us to find hypothesis functions that:
 - Use multiple features (input variables), e.g., $H(x) = w_0 + w_1x^{(1)} + w_2x^{(2)}$.
 - Are non-linear in the features, e.g., $H(x) = w_0 + w_1x + w_2x^2$.
- Let's see if we can put what we learned last week to use.

Simple linear regression, revisited



- Model: $H(x) = w_0 + w_1x$.
- Loss function: $(y_i - H(x_i))^2$.
- To find w_0^* and w_1^* , we minimized empirical risk, i.e. average loss:

$$R_{\text{sq}}(H) = \frac{1}{n} \sum_{i=1}^n (y_i - H(x_i))^2$$

Handwritten annotations: "observed" (orange) points to y_i , "prediction" (blue) points to $H(x_i)$, "loss" (blue) points to the squared difference, and "averaging" (blue) points to the $1/n$ factor.

- Observation: $R_{\text{sq}}(w_0, w_1)$ kind of looks like the formula for the norm of a vector,

$$\|\vec{v}\| = \sqrt{v_1^2 + v_2^2 + \dots + v_n^2}.$$

Regression and linear algebra

Let's define a few new terms:

- The **observation vector** is the vector $\vec{y} \in \mathbb{R}^n$. This is the vector of observed values.
- The **hypothesis vector** is the vector $\vec{h} \in \mathbb{R}^n$ with components $H(x_i)$. This is the vector of predicted values.
- The **error vector** is the vector $\vec{e} \in \mathbb{R}^n$ with components:

$$e_i = y_i - H(x_i)$$

$$\vec{e} = \vec{y} - \vec{h}$$

This is the vector of signed errors.

$$\vec{y} = \begin{bmatrix} 35 \text{ minutes} \\ 72 \text{ minutes} \\ 27 \text{ minutes} \\ \vdots \end{bmatrix}$$

$$\vec{h} = \begin{bmatrix} H(x_1) \\ H(x_2) \\ \vdots \\ H(x_n) \end{bmatrix} = \begin{bmatrix} 41 \text{ minutes} \\ 70 \text{ minutes} \\ \vdots \end{bmatrix}$$

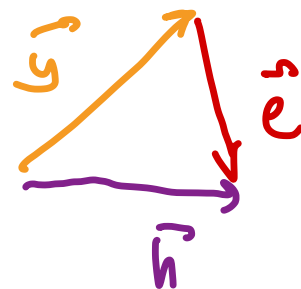
$$\vec{e} = \begin{bmatrix} e_1 = y_1 - H(x_1) \\ e_2 = y_2 - H(x_2) \\ \vdots \end{bmatrix}$$

Regression and linear algebra

Let's define a few new terms:

- The **observation vector** is the vector $\vec{y} \in \mathbb{R}^n$. This is the vector of observed values.
- The **hypothesis vector** is the vector $\vec{h} \in \mathbb{R}^n$ with components $H(x_i)$. This is the vector of predicted values.
- The **error vector** is the vector $\vec{e} \in \mathbb{R}^n$ with components: $e_i = y_i - H(x_i)$
- **Key idea:** We can rewrite the mean squared error of H as:

$$R_{\text{sq}}(H) = \frac{1}{n} \sum_{i=1}^n (y_i - H(x_i))^2 = \frac{1}{n} \sum_{i=1}^n e_i^2 = \frac{1}{n} \|\vec{e}\|^2 = \frac{1}{n} \|\vec{y} - \vec{h}\|^2$$



The hypothesis vector

- The hypothesis vector is the vector $\vec{h} \in \mathbb{R}^n$ with components $H(x_i)$. This is the vector of predicted values.
- For the linear hypothesis function $H(x) = w_0 + w_1x$, the hypothesis vector can be written:

$$\vec{h} = \begin{bmatrix} w_0 + w_1x_1 \\ w_0 + w_1x_2 \\ \vdots \\ w_0 + w_1x_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \begin{bmatrix} w_0 \\ w_1 \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} w_0 + \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix} w_1$$

every element (with arrow pointing to the first element of the hypothesis vector)

$h_i = H(x_i) = w_0 + w_1x_i$

X -design matrix (with arrow pointing to the matrix in the second term)

model parameters (with arrow pointing to the vector in the second term)

$\vec{1}$ all-ones (with arrow pointing to the vector in the first term)

Rewriting the mean squared error

- Define the design matrix $X \in \mathbb{R}^{n \times 2}$ as:

$$X = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix}$$

- Define the parameter vector $\vec{w} \in \mathbb{R}^2$ to be $\vec{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$.

\nearrow slope
 \searrow intercept

- Then, $\vec{h} = X\vec{w}$, so the mean squared error becomes:

$$R_{\text{sq}}(H) = \frac{1}{n} \|\vec{y} - \vec{h}\|^2 \implies \boxed{R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2}$$

Minimizing mean squared error, again

- To find the optimal model parameters for simple linear regression, w_0^* and w_1^* , we previously minimized:

$$R_{\text{sq}}(w_0, w_1) = \frac{1}{n} \sum_{i=1}^n (y_i - (w_0 + w_1 x_i))^2$$

- Now that we've reframed the simple linear regression problem in terms of linear algebra, we can find w_0^* and w_1^* by finding the $\vec{w}^* = [w_0^* \quad w_1^*]^T$ that minimizes:

$$R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2$$

- Do we already know the \vec{w}^* that minimizes $R_{\text{sq}}(\vec{w})$?

An optimization problem we've seen before

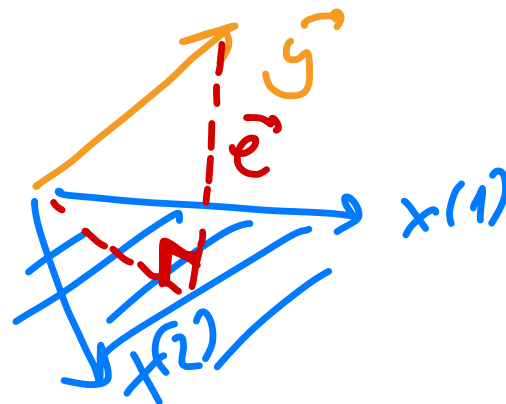
- The optimal parameter vector, $\vec{w}^* = [w_0^* \quad w_1^*]^T$, is the one that minimizes:

$$R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2 = \frac{1}{n} \|\vec{e}\|^2$$

- The minimizer of $\|\vec{e}\|$ is the same as the minimizer of $R_{\text{sq}}(\vec{w})$!

$$\vec{w}^* = \arg \min_{\vec{w}} R_{\text{sq}} = \arg \min_{\vec{w}} \|\vec{e}\|$$

- Last week we found that the vector in the span of the columns of X that is closest to \vec{y} is the vector $X\vec{w}$ such that $\|\vec{e}\| = \|\vec{y} - X\vec{w}\|$ is minimized.



The modeling recipe

1. Choose a model.

$$H(x) = [1 \quad x]^T \vec{w} = w_0 + w_1 x \quad \text{SLR}$$

2. Choose a loss function.

squared loss

$$e^2 = (y - [1 \quad x]^T w)^2$$

3. Minimize average loss to find optimal model parameters.

$$\vec{w}^* = \arg \min_{\vec{w}} R_{\text{sq}}(\vec{w}) = \arg \min_{\vec{w}} \left\{ \frac{1}{n} \|\vec{y} - X\vec{w}\|^2 \right\} = \arg \min_{\vec{w}} \left\{ \frac{1}{n} \|\vec{e}\|^2 \right\}$$

$$\vec{a} \cdot \vec{b} = 0 \iff \vec{a}, \vec{b} \text{ are orthogonal}$$

An optimization problem we've seen before

- Key idea: Find $\vec{w} \in \mathbb{R}^d$ such that the **error vector**, $\vec{e} = \vec{y} - X\vec{w}$, is **orthogonal** to the columns of X .

- Why? Because this will make the **error vector** as short as possible.

- The \vec{w}^* that accomplishes this satisfies:

$$X^T \vec{e} = 0$$

$$X^T (\vec{y} - X\vec{w}) = 0$$

- Why? Because $X^T \vec{e}$ contains the **dot products** of each column in X with \vec{e} . If these are all 0, then \vec{e} is **orthogonal to every column of X** !

$$X = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix}$$

$A\vec{v}$ -
dot product of
 \vec{v} with the rows
of A

$$X^T \vec{e} = \begin{bmatrix} -\vec{1}^T - \\ -\vec{x}^T - \end{bmatrix} \vec{e} = \begin{bmatrix} \vec{1}^T \vec{e} \\ \vec{x}^T \vec{e} \end{bmatrix}$$

$$\begin{bmatrix} 1 & 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix} \vec{e} = 0$$

The normal equations

- Key idea: Find $\vec{w} \in \mathbb{R}^d$ such that the **error vector**, $\vec{e} = \vec{y} - X\vec{w}$, is **orthogonal** to the columns of X .

- The \vec{w}^* that accomplishes this satisfies:
- Assuming $X^T X$ is invertible, this is the vector:

$$X^T \vec{e} = 0$$

$$X^T (\vec{y} - X\vec{w}^*) = 0$$

$$X^T \vec{y} - X^T X \vec{w}^* = 0$$

$$\vec{w}^* = (X^T X)^{-1} X^T \vec{y} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$

- The normal equations:

$$\implies X^T X \vec{w}^* = X^T \vec{y}$$

~~$$(X^T X)^{-1} X^T X \vec{w}^* = (X^T X)^{-1} X^T \vec{y}$$~~

- This is a big assumption, because it requires $X^T X$ to be **full rank**. *all columns are linearly independent*
- If $X^T X$ is not full rank, then there are infinitely many solutions to the normal equations.

X is full rank

An optimization problem, solved

(no calculus)

- We just used linear algebra to solve an optimization problem.
- Specifically, the function we minimized is:

$$\text{error}(\vec{w}) = \|\vec{y} - X\vec{w}\|$$

- The input, \vec{w}^* , to $\text{error}(\vec{w})$ that minimizes it is one that satisfies the normal equations:

$$X^T X \vec{w}^* = X^T \vec{y}$$

If $X^T X$ is invertible, then the unique solution is:

$$\vec{w}^* = (X^T X)^{-1} X^T \vec{y}$$

- Key idea: $\vec{w}^* = (X^T X)^{-1} X^T \vec{y}$ also minimizes $R_{\text{sq}}(\vec{w})$!
- We're going to use this frequently!

Alternative solution

- Our goal is to find the vector \vec{w} that minimize mean squared error:

$$R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2$$

- Strategy: calculus
- Problem: This is a function *of a vector*. What does it even mean to take the derivative of $R_{\text{sq}}(\vec{w})$ with respect to a vector \vec{w} ?

A function of a vector

- **Solution:** A function *of a vector* is really just a function *of multiple variables*, which are the components of the vector. In other words,

$$R_{\text{sq}}(\vec{w}) = R_{\text{sq}}(w_0, w_1, \dots, w_d) \in \mathbb{R}$$

$$\vec{w} = \begin{bmatrix} w_1 \\ w_2 \\ \vdots \\ w_d \end{bmatrix} \in \mathbb{R}^d$$

where w_0, w_1, \dots, w_d are the entries of the vector \vec{w} .

In our case, \vec{w} has just two components, w_0 and w_1 . We'll be more general since we eventually want to use prediction rules with even more parameters.

- We know how to deal with derivatives of multivariable functions: the gradient!

The gradient with respect to a vector

- The gradient of $R_{\text{sq}}(\vec{w})$ with respect to \vec{w} is the vector of partial derivatives:

$$\begin{bmatrix} \frac{\partial R_{\text{sq}}}{\partial w_0} \\ \frac{\partial R_{\text{sq}}}{\partial w_1} \end{bmatrix} \in \mathbb{R}^2 = \nabla_{\vec{w}} R_{\text{sq}}(\vec{w}) = \frac{dR_{\text{sq}}}{d\vec{w}} = \begin{bmatrix} \frac{\partial R_{\text{sq}}}{\partial w_0} \\ \frac{\partial R_{\text{sq}}}{\partial w_1} \\ \vdots \\ \frac{\partial R_{\text{sq}}}{\partial w_d} \end{bmatrix} \in \mathbb{R}^d$$

where w_0, w_1, \dots, w_d are the entries of the vector \vec{w} .

Goal

- We want to minimize the mean squared error: as a function of vector \vec{w}

$$R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2$$

- Strategy:

1. Compute the gradient of $R_{\text{sq}}(\vec{w})$.

2. Set it to zero and solve for \vec{w} .

- The result is the optimal parameter vector \vec{w}^* .

- Let's start by rewriting the mean squared error in a way that will make it easier to compute its gradient.

Question 🤔

Answer at q.dsc40a.com

Which of the following is equivalent to $R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2$?

A) $\frac{1}{n} (\vec{y} - X\vec{w}) \cdot (X\vec{w} - y)$

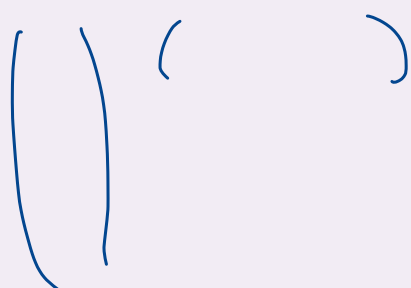
B) $\frac{1}{n} \sqrt{(\vec{y} - X\vec{w}) \cdot (y - X\vec{w})} = \frac{1}{n} \|\vec{e}\|$

C) $\frac{1}{n} (\vec{y} - X\vec{w})^T (y - X\vec{w})$

D) $\frac{1}{n} (\vec{y} - X\vec{w})(y - X\vec{w})^T$

$$\frac{1}{n} \|\vec{e}\|^2 = \frac{1}{n} \vec{e} \cdot \vec{e} = \frac{1}{n} \vec{e}^T \vec{e}$$

$$\begin{aligned} &= \frac{1}{n} (\vec{y} - X\vec{w})^T (\vec{y} - X\vec{w}) \\ &= \frac{1}{n} (\vec{y} - X\vec{w}) \cdot (y - X\vec{w}) \end{aligned}$$



Rewriting mean squared error

Reminder:

$$(AB)^T = B^T A^T$$

$$A(BC) = (AB)C$$

$$R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2 =$$

$$= \frac{1}{n} (\vec{y} - X\vec{w})^T (\vec{y} - X\vec{w})$$

$$= \frac{1}{n} (\vec{y}^T - (X\vec{w})^T) (\vec{y} - X\vec{w})$$

$$= \frac{1}{n} (\vec{y}^T - \vec{w}^T X^T) (\vec{y} - X\vec{w})$$

$$= \frac{1}{n} (\vec{y}^T \vec{y} - \vec{y}^T (X\vec{w}) - (\vec{w}^T X^T) \vec{y} + \vec{w}^T X^T X \vec{w}) =$$

$$\vec{y} \cdot (X\vec{w}) = (X^T \vec{y}) \cdot \vec{w} = \vec{w} \cdot (X^T \vec{y})$$

$$= \frac{1}{n} \left(\vec{y}^T \vec{y} - (X^T \vec{y}) \cdot \vec{w} - \vec{w} \cdot (X^T \vec{y}) + \vec{w}^T X^T X \vec{w} \right)$$

$$= \frac{1}{n} \left(\vec{y}^T \vec{y} - 2(X^T \vec{y}) \cdot \vec{w} + \vec{w}^T X^T X \vec{w} \right)$$

$$\vec{a} \cdot \vec{b} = \vec{b} \cdot \vec{a}$$

Compute the gradient

$$\begin{aligned}\frac{dR_{\text{sq}}}{d\vec{w}} &= \frac{d}{d\vec{w}} \left(\frac{1}{n} (\vec{y} \cdot \vec{y} - 2\mathbf{X}^T \vec{y} \cdot \vec{w} + \vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) \right) \\ &= \frac{1}{n} \left(\frac{d}{d\vec{w}} (\vec{y} \cdot \vec{y}) - \frac{d}{d\vec{w}} (2\mathbf{X}^T \vec{y} \cdot \vec{w}) + \frac{d}{d\vec{w}} (\vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) \right)\end{aligned}$$

Question 🤔

Answer at q.dsc40a.com

Which of the following is $\frac{d}{d\vec{w}} (\vec{y} \cdot \vec{y})$?

A. $\vec{y} \cdot \vec{y}$

B. $2\vec{y}$

C. 1

D. 0

\vec{y} doesn't depend on \vec{w}

Compute the gradient

$$\begin{aligned} \frac{dR_{\text{sq}}}{d\vec{w}} &= \frac{d}{d\vec{w}} \left(\frac{1}{n} (\vec{y} \cdot \vec{y} - 2\mathbf{X}^T \vec{y} \cdot \vec{w} + \vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) \right) \\ &= \frac{1}{n} \left(\frac{d}{d\vec{w}} (\vec{y} \cdot \vec{y}) - \frac{d}{d\vec{w}} (2\mathbf{X}^T \vec{y} \cdot \vec{w}) + \frac{d}{d\vec{w}} (\vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) \right) \end{aligned}$$

○
 $2\mathbf{X}^T \vec{y}$
 $2\mathbf{X}^T \mathbf{X} \vec{w}$

- $\frac{d}{d\vec{w}} (\vec{y} \cdot \vec{y}) = 0.$
 - Why? \vec{y} is a constant with respect to \vec{w} .
- $\frac{d}{d\vec{w}} (\vec{2}\mathbf{X}^T \vec{y} \cdot \vec{w}) = 2\mathbf{X}^T \vec{y}.$
 - Why? In groupwork today you will show $\frac{d}{d\vec{x}} \vec{a} \cdot \vec{x} = \vec{a}.$
- $\frac{d}{d\vec{w}} (\vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) = 2\mathbf{X}^T \mathbf{X} \vec{w}.$
 - Why? You will prove in homework 4.

Compute the gradient

$$\begin{aligned}\frac{dR_{\text{sq}}}{d\vec{w}} &= \frac{d}{d\vec{w}} \left(\frac{1}{n} (\vec{y} \cdot \vec{y} - 2\mathbf{X}^T \vec{y} \cdot \vec{w} + \vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) \right) \\ &= \frac{1}{n} \left(\frac{d}{d\vec{w}} (\vec{y} \cdot \vec{y}) - \frac{d}{d\vec{w}} (2\mathbf{X}^T \vec{y} \cdot \vec{w}) + \frac{d}{d\vec{w}} (\vec{w}^T \mathbf{X}^T \mathbf{X} \vec{w}) \right) \\ &= \frac{1}{n} (-2\mathbf{X}^T \vec{y} + 2\mathbf{X}^T \mathbf{X} \vec{w})\end{aligned}$$

The normal equations (again)

- To minimize $R_{\text{sq}}(\vec{w})$, set its gradient to zero and solve for \vec{w} :

$$\begin{aligned} -2X^T \vec{y} + 2X^T X \vec{w} &= 0 \\ \implies X^T X \vec{w} &= X^T \vec{y} \end{aligned}$$

- We have seen this system of equations in matrix form before: the **normal equations**.
- If $X^T X$ is invertible, the solution is

$$\vec{w}^* = (X^T X)^{-1} X^T \vec{y}$$

The optimal parameter vector, \vec{w}^*

- To find the optimal model parameters for simple linear regression, w_0^* and w_1^* , we previously minimized $R_{\text{sq}}(w_0, w_1) = \frac{1}{n} \sum_{i=1}^n (y_i - (w_0 + w_1 x_i))^2$.

- We found, using calculus, that:

- $$w_1^* = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = r \frac{\sigma_y}{\sigma_x}.$$

- $$w_0^* = \bar{y} - w_1^* \bar{x}.$$

- Another way of finding optimal model parameters for simple linear regression is to find the \vec{w}^* that minimizes $R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2$.

- The minimizer, if $X^T X$ is invertible, is the vector
$$\vec{w}^* = (X^T X)^{-1} X^T \vec{y}.$$

- These formulas are equivalent!

Summary: Regression and linear algebra (Solution 1)

- Define the design matrix $X \in \mathbb{R}^{n \times 2}$, observation vector $\vec{y} \in \mathbb{R}^n$, and parameter vector $\vec{w} \in \mathbb{R}^2$ as:

$$X = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \quad \vec{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} \quad \vec{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$

- How do we make the hypothesis vector, $\vec{h} = X\vec{w}$, as close to \vec{y} as possible? Use the parameter vector \vec{w}^* :

$$\vec{w}^* = (X^T X)^{-1} X^T \vec{y}$$

- We chose \vec{w}^* so that $\vec{h}^* = X\vec{w}^*$ is the projection of \vec{y} onto the span of the columns of the design matrix, X and minimized the length of the projection error $\|\vec{e}\| = \|\vec{y} - X\vec{w}^*\|$.

Summary: Regression and linear algebra (Solution 2)

- Define the design matrix $X \in \mathbb{R}^{n \times 2}$, observation vector $\vec{y} \in \mathbb{R}^n$, and parameter vector $\vec{w} \in \mathbb{R}^2$ as:

$$X = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \quad \vec{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} \quad \vec{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$

- How do we minimize the mean squared error $R_{\text{sq}}(\vec{w}) = \frac{1}{n} \|\vec{y} - X\vec{w}\|^2$? Using calculus the optimal parameter vector \vec{w}^* is:

$$\vec{w}^* = (X^T X)^{-1} X^T \vec{y}$$

Roadmap

- Next class, we'll present a more general framing of the multiple linear regression model, that uses d features instead of just two.
- We'll also look at how we can **engineer** new features using existing features.
 - e.g. How can we fit a hypothesis function of the form
$$H(x) = w_0 + w_1x + w_2x^2?$$